

New Robust Beta Regression Estimation to Overcome the Effect of High Leverage Points

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ABSTRACT

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Beta regression models are widely used to model continuous data with a unit interval (0,1), such as ratios, fractions, and rates. The maximum likelihood method is typically used to estimate regression coefficients in beta regression models. However, the maximum likelihood estimator is highly sensitive to outliers. Several studies have proposed methods to address this problem, such as the M-Huber, S-estimator, LMS, and LTS estimators. However, these methods suffer from the problem of high leverage points (HLPs) in the independent variables. The GM-estimator is one of the methods that address this problem. This study examines beta regression analysis, focusing on the effect of high leverage points (HLPs) on parameter estimates. Monte Carlo simulations and real data are conducted to evaluate and compare the performance of the proposed robust method, Generalized M-Beta Regression (GMBR), with the existing beta regression method estimated using Maximum Likelihood (MLBr).

Keywords: Beta regression; Outliers; High leverage points (HLPs); Generalized M, Robust estimation.

1- Introduction

Regression analysis is one of the most widely used statistical methods in various fields. It describes the relationship between a response variable and one or more independent variables, with the aim of determining the form of their association through a mathematical model. When the response variable (y) is expressed as fractions or percentages -- that is, when its values are limited to the interval (0, 1), beta regression is typically used. Beta regression is a well-known statistical model that finds applications in the natural sciences, medicine, finance, economics, environment, hydrology, psychology, and many other disciplines.

Maximum likelihood estimations (MLE) are commonly used to estimate the parameters of a beta regression model. (Ferrari & Cribari-Neto, 2004) proposed the beta regression model, and several studies have been conducted on different aspects of beta regression. For example, (Vasconcellos & Cribari-Neto, 2005) investigated the behavior of MLEs in a beta regression model where the distribution parameters are nonlinear functions of linear combinations of explanatory variables with unknown coefficients. (Ospina et al., 2006) conducted a study that involved point and interval optimization of a beta regression model. (Bayes et al., 2012a) presented a new regression model for proportions by incorporating the rectangular beta distribution proposed by (Bayes et al., 2012b) Hahn (2008). (Espinheira et al., 2015) studied separation measures in beta regression models, and (Bayer & Cribari-Neto, 2017) presented a study titled "Model selection criteria in beta regression with variable dispersion." (Espinheira et al., 2019) presented a study on model selection criteria for beta regression, and (Ghosh, 2019) conducted a study on developing a robust inference procedure for beta regression model. Additional studies have contributed to the development of beta regression methodologies.

(Karlsson et al., 2020) proposed Liu shrinkage estimators for beta regression models, whereas (Abonazel, Dawoud, et al., 2022) proposed the Dawood-Kabria estimator for beta regression models. (Zhou & Huang, 2022) presented a study involving Bayesian empirical regression for limited responses with unknown support. (Abonazel & Taha, 2023) studied beta ridge regression estimators. (Ospina et al., 2023) proposed a robust semi-parametric inference for two-stage production models using a beta regression approach. (Wilcox, 2011) conducted a comparative study on robust estimation of a beta regression model in the presence of outliers. (Ribeiro & Ferrari, 2023) investigated robust estimation in beta regression using the Lq maximum likelihood method (Maluf et al., 2024) applied robust beta regression using logit transformation, and (Heng & Lange, 2025) conducted a study involving a preliminary estimate of the proportion of outliers in robust regression.

The success or failure of the estimator depends on the method used to estimate the model parameters. The ordinary least squares (OLS) method is known to be very sensitive to outliers as it has a breaking point close to (0), and the inference based on the maximum likelihood estimator lacks power when outliers are present. This sensitivity can lead to severe bias and misleading conclusions. Outliers represent one of the oldest and most important challenges in statistics. (Rosseeuw & Van Zomeren, 1990) classified outliers into vertical outliers that appear in the response variable (y), leverage points that appear in the independent variables (x), and outliers in both directions. When diagnosing the problem of the presence of outliers, many researchers treated the outliers by removing them and analyzing the remaining observations. Other studies have emphasized the importance of retaining these data, because any lack of observations can lead to the loss of valuable information and negatively affect the accuracy of the results. Instead, robust methods for dealing with outliers have been used, ensuring that the estimators and their associated test statistics are relatively insensitive to the presence of outliers. These robust methods, which include the M-Huber, S-estimator, LMS, and LTS estimators (Lim & Midi, 2016), generally provide more efficient results than traditional methods. However, even these robust methods can be affected by high leverage points (HLPs) in the independent variables. In this study, we modeled the data using beta regression and used maximum likelihood estimation to estimate the model parameters. To address the problem of outliers, we used generalized M-estimators (GM estimators) that can be powerful tools for mitigating the effect of high leverage points (HLPs) in the independent variables.

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eta Regression Model.

Practitioners typically use linear regression models to examine the relationship between a dependent variable and selected explanatory variables, as well as to analyze their effects. However, this approach is unsuitable when the dependent variable is constrained within the interval (0, 1), since it may lead to predicted values that exceed these boundaries. Consequently, inferences based on the normality assumption can be misleading.

The beta regression model was first introduced by (Ferrari & Cribari-Neto, 2004), linking the mean of the dependent variable to a set of linear predictors through a monotonic and differentiable link function. This model includes a precision parameter, the inverse of which is referred to as the dispersion scale. In the basic form of the beta regression model, the precision parameter is assumed to be constant across all observations. However, some studies—such as those by (Smithson & Verkuilen, 2006) and (Algama et al., 2023) - have indicated that the precision parameter may vary across observations in certain cases.

Let y is a continuous random that follows a beta distribution with the following probability density function;

$$f(y; \mu, \lambda) = \frac{\Gamma(\lambda)}{\Gamma(\mu\lambda)\Gamma((1-\mu)\lambda)} y^{(\mu\lambda)-1} (1-y)^{(1-\mu)\lambda-1}, \quad 0 < y < 1; 0 < \mu < 1; \lambda > 0 \quad (1)$$

where $\Gamma(\cdot)$ is the gamma function, and λ is the precision parameter that can be written as in (Bayer & Cribari-Neto, 2017):

$$\lambda = \frac{1-\sigma^2}{\sigma^2} \quad (2)$$

The mean and variance of the beta probability distribution are: $E(y) = \mu$, $var(y) = \mu(1 - \mu)\sigma^2$. Using the logit link function, the model allows μ_i depending on covariates as follows

$$g(\mu_i) = \log\left(\frac{\mu_i}{1-\mu_i}\right) = x_i' \beta = \eta_i \quad (3)$$

where $g(\cdot)$ be a monotonic differentiable link function used to relate the systematic component with the random component, $\beta = (\beta_1, \dots, \beta_K)'$ is a $(k \times 1)$ vector of unknown parameters, $x_i = (x_{i1}, \dots, x_{ik})'$ is the vector of k regression, and η_i is the linear predictor. A particularly interesting specification of the model is obtained when the logit link function is used. In the case, the mean of y_i can be written as (El-Raoof et al., 2023) by:

$$\mu_i = \frac{e^{x_i' \beta}}{1 + e^{x_i' \beta}} \quad (4)$$

Where μ_i is the mean response function. since η depends on β and the mean response μ is a function of η , the mean $\mu_1, \mu_2, \dots, \mu_n$ are function of β (Abonazel, Algamal, et al., 2022).

Estimation of the beta regression parameters is done by using the ML method Espinheira et al., (2008) The log-likelihood function of the beta regression model is given by:

$$L(\mu_i, \lambda; y_i) = \sum_{i=1}^n \{ \log \Gamma(\lambda) - \log \Gamma(\mu_i(\lambda)) - \log \Gamma((1 - \mu_i)(\lambda)) + (\mu_i(\lambda) - 1) \log(y_i) + ((1 - \mu_i)(\lambda) - 1) \log(1 - y_i) \} \quad (5)$$

Differentiating the log-likelihood in Eq. (5) with respect to β gives us the score function for β which is given by:

$$U(\beta) = \lambda x' F(y^* - \mu^*) \quad (6)$$

Where $F = \text{diag}\left(\frac{1}{g'(\mu_1)}, \dots, \frac{1}{g'(\mu_n)}\right)$, $y^* = (y_1^*, \dots, y_n^*)'$ and $y_i^* = \log\left(\frac{y_i}{1-y_i}\right)$

$\mu^* = (\mu_1^*, \dots, \mu_n^*)'$ and $\mu_i^* = \varphi(\mu_i \lambda) - \varphi((1 - \mu_i)\lambda)$ such that $\varphi(\cdot)$ denoting the digamma function. The iterative reweighted least-squares (IWLS) algorithm or Fisher scoring algorithm used for estimating β (Espinheira et al., 2015). The form of this algorithm can be written as :

$$\beta^{(t+1)} = \beta^{(t)} + (I_{\beta\beta}^{(t)})^{-1} U_{\beta}^{(t)}(\beta)$$

Where $U_{\beta}^{(r)}$ is the score function defined in Eq. (6), and $I_{\beta\beta}^{(r)}$ is the information matrix for β , see (Espinheira et al., 2019) and (Algamal et al., 2023) for more details. The initial value of β can be obtained by the least squares method, while the initial value for each precision parameter equals

$$\hat{\lambda}_i = \frac{\hat{\mu}_i(1-\hat{\mu}_i)}{\hat{\sigma}_i^2} \quad (7)$$

Where $\hat{\mu}$ and σ_i^2 values are obtained from linear regression .Given $r = 0,12, \dots$ is the number of iterations that are performed, convergence occurs when the difference between successive estimates becomes smaller than a given small constant .At the final step, the ML estimator of β is obtained as (Abonazel, Dawoud, et al., 2022):

$$\hat{\beta}_{GM} = (x' \hat{W} x)^{-1} x' \hat{W} \hat{z} \quad (8)$$

Where X is an $n \times p$ matrix of regressors, $\hat{z} = \hat{\eta} + \hat{W}^{-1} \hat{F} (y^* - \mu^*)$, and $\hat{W} = \text{diag}(\hat{w}_1, \dots \dots \hat{w}_n)$

$$\hat{w}_i = \left(\frac{1-\hat{\sigma}^2}{\hat{\sigma}^2} \right) \left\{ \hat{\phi} \left(\frac{\hat{\mu}_i(1-\hat{\sigma}^2)}{\hat{\sigma}^2} \right) + \phi' \left(\frac{(1-\hat{\mu}_i)(1-\hat{\sigma}^2)}{\hat{\sigma}^2} \right) \right\} \frac{1}{\{g'(\hat{\mu}_i)\}^2}.$$

4- Generalized M-Estimator method (GM-estimator)

The generalized M-estimator (GM-estimator) is an extension of the M-estimator method introduced by Peter J. Huber in 1964. This method aims to provide robust estimates of parameters in statistical models, especially when there are extreme values or non-normal distributions of the data (Cheng & Van Ness, 1992). It was proposed by Schwepes, as described by Hill and Paul (1977), and was developed by several researchers, most notably Frank and Hampel in the 1970s and 1980s. The basic idea behind GM estimators is to minimize the effect of high leverage points, which is the main problem with M estimators (Hampel et al., n.d.) and (Midi et al., 2021). GM estimators can deal with this problem by taking advantage of some weight functions that minimize the presence of high leverage points. The GM estimator is the solution to the following equation (Wilcox, 2011).

$$\sum_{i=1}^n \tau_i \psi \left(\frac{(y_i - x_i' \hat{\beta})}{s \tau_i} \right) x_i \quad (9)$$

where τ_i is an initial weight function used to minimize the effects of leverage points, the GM-estimator in convergence can be written as

$$\hat{\beta}_{GM} = (x' \hat{W} x)^{-1} x' \hat{W} y \quad (10)$$

where W is a diagonal weight matrix with elements w_i defined as

$$w_i = \frac{\psi[(y_i - x_i' \hat{\beta}_{GML}) / \tau_i s]}{(y_i - x_i' \hat{\beta}_{GML}) / \tau_i s} \quad (11)$$

The initial weights of GM-estimators that minimize the effect of leverage points in (9) are computed based on the hat matrix values h_{ii} as (Wilcox, 2011)

$$\tau_i = \sqrt{(1 - h_{ii})} \quad (12)$$

(h_{ii}) is found through a (hat matrix) which is a weight matrix, this matrix is used to identify the rows of existing observations (x) that contain extreme values, and is known as regression analysis to detect the presence of (HLPoints), and these rows can be written according to the following formula (Iglewicz & Hoaglin, 1993)

$$H = x'(x'x)^{-1}x \quad (13)$$

The elements of the main diagonal of the (Hat Matrix) H are written as follows (Rousseeuw & Leroy, 2003):

$$h_{ii} = x'(x'x)^{-1}x_i \quad (14)$$

These elements have desirable properties as their values range between (1 0). Also, the h_{ii} set equals p and its containment index of (HLPpoints) for case j is a measure of the distance between the values of (x) for case j and the average values of (x) for all cases (n) . Therefore, the large values (h_{ii}) indicate the containment index of case j , and thus it is far from the center of all observations of the variable (x) . The average of this matrix can be written in the following formula (PJ & Leroy, 1987)

$$H_n^p = \frac{\sum_1^n h_{ii}}{n} =$$

p : the number of independent variables, n : the sample size.

Both (Hoaglin & Welsch, 1978) suggested that the threshold value should be when there are (HLPpoints) in one variable and not another when the value is $H_{tt} > \frac{2p}{n}$, and there is also a rule of three times the mean that was presented by (Velleman & Welsch, 1981) in order to diagnose (HLPpoints) when it is $H_{tt} > \frac{3p}{n}$, and Huber, 2004, presented another cut-off point as a three-point interval for the data ($0.2 < H_{tt} < 0.5$), and accordingly the data values should be avoided ($H_{tt} > 0.5$), Kutner et al., 2005, noticed that the presence of a gap between the data for the stupidest cases and the abnormal values in an unusual way is conclusive evidence of the presence of abnormal value

6- Algorithm.

GM method is one of the important methods for treating outliers and leverage points as it can be used to reduce the effect of leverage points in independent variables (x) when estimating the parameters of beta regression model. An algorithm for GM-estimator can be written in the following steps:

Step 1: Choosing an initial estimates $\beta^{(0)}$ from beta regression.

Step 2: For each iteration r compute $\mu_i^{(t-1)} = \text{logistic}(x_i \beta^{(t-1)}) = \frac{\exp(x_i \beta^{(t-1)})}{1 + \exp(x_i \beta^{(t-1)})}$.

Step 3: Compute the residuals $r_i^{(t-1)} = y_i - \mu_i^{(t-1)}$ and scale $\hat{\sigma}^{(t-1)} = 1.4826 * \text{median of largest}(n - p) \text{ of the } |r_i^{(t-1)}|$, and then compute the standardized residuals $u_i^{(t-1)} = \frac{r_i^{(t-1)}}{\hat{\sigma}^{(t-1)}}$.

Step 4: Using the following form to compute the weight $\tau_i = \sqrt{(1 - h_{ii})}$.

Step 5: Based on Huber weight function standardized residuals u_i use to compute the robust weights using the form $w_i^{(t-1)} = \frac{\psi\left(\frac{(y_i - \mu_i^{(t-1)})}{\tau_i \hat{\sigma}}\right)}{\left(\frac{(y_i - \mu_i^{(t-1)})}{\tau_i \hat{\sigma}}\right)}$.

Step 6: compute $\hat{\beta}^t$ using the form $\hat{\beta}^t = (X'W^{(t-1)}X)^{-1}X'W^{(t-1)}Y$.

Step 7: Steps (2-6) are repeated until convergence.

Simulation study

This section presents a Monte Carlo simulation to assess and compare the performance of the proposed robust method Generalized-M for Beta regression (GMBR) with the existing Beta regression method estimated using maximum likelihood (MLBR). In this study, we also propose a the

performance criterion that considered by (Abonazel & Taha, 2023) , referred to as mean squares error for $(\hat{\beta})$ ($MSE(\hat{\beta})$) and mean absolute error for $(\hat{\beta})$ ($MAE(\hat{\beta})$), it can calculated as follows:

$$MSE(\hat{\beta}) = \frac{1}{T} \sum_{t=1}^T (\hat{\beta}_t - \beta^{true})' (\hat{\beta}_t - \beta^{true}), \quad MAE((\hat{\beta})) = \frac{1}{T} \sum_{t=1}^T |\hat{\beta}_t - \beta^{true}|$$

In this study, we generated the explanatory variables from a multivariate normal distribution $MN(0,1)$ with four different sample sizes ($n = 30, 50, 100$, and 250). Additionally, we assume two scenarios for the coefficients and the number of explanatory variables:

- $(\beta_1 = 1, \beta_j = 0.25 \text{ for all } j = 2, \dots, 5)$,
- $(\beta_1 = 2, \beta_j = 0.50 \text{ for all } j = 2, \dots, 7)$

where β_1 is the intercept, the number of explanatory variables p set as 4 and 6 . Three levels of

contamination for the explanatory variables are considered ($\delta = 5\%, 10\%, 20\%$) and the

precision parameter is set to as $\lambda=2$ and 4. The response variable will be generated as $y_i \sim \text{Beta}(\mu_i, \lambda)$, and the experiment will be repeated 500 times ($T=500$). Based on the package “

betareg” introduced by (Cribari-Neto & Zeileis, 2010), the R language was employed to conduct our simulation study.

Table(1): the $MSE(\hat{\beta})$ and $MAE(\hat{\beta})$ at Three Different Contamination Levels and different Sample Sizes in the First Scenario When the Precision Parameter is $\lambda=2$.

Con t.	Samp le	metho ds	$MSE(\hat{\beta})$					$MAE(\hat{\beta})$				
			$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$	$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$
5%	30	BR	0.955 5	0.48 46	1.428 4	0.96 29	1.581 4	1.385 6	1.065 1	1.986 8	1.633 7	2.16 01
		GMBR	0.94 21	0.213 2	1.143 8	0.710 3	1.313 7	1.707 5	0.837 7	1.739 1	1.432 4	1.927 3
	50	BR	0.90 65	0.34 02	1.266 6	0.84 99	1.45 09	1.182	0.691 1	1.583	1.316 3	1.79 89
		GMBR	0.831	0.124	1.053 8	0.62 73	1.22 68	1.325 6	0.49 64	1.396 2	1.103 3	1.58 98
	100	BR	0.88 35	0.265 4	1.194 4	0.76 61	1.36 55	1.057 6	0.451 6	1.348 3	1.051 7	1.53 81
		GMBR	0.751 4	0.06 26	0.99 26	0.56 29	1.162 4	1.044 8	0.25 45	1.149 9	0.85 05	1.337 6
	250	BR	0.774	0.125	1.055 9	0.62 62	1.22 49	0.89 53	0.199 3	1.101 2	0.80 16	1.287 2
		GMBR										

		GMBR	0.70 26	0.02 5	0.95 51	0.525 1	1.124 9	0.87 41	0.187 7	1.00 02	0.70 02	1.187 8
10%	30	BR	0.96 06	0.40 02	1.405 6	0.96 78	1.556 6	1.416 7	0.94 37	1.981 2	1.666 3	2.13 49
		GMBR	0.94 54	0.199 1	1.143 5	0.713 3	1.312 1	1.713 5	0.811 4	1.740 3	1.444 3	1.925 6
	50	BR	0.90 47	0.33 27	1.275 7	0.83 66	1.427 8	1.173 7	0.69 36	1.626 4	1.296 2	1.774 2
		GMBR	0.83 41	0.124 5	1.057 7	0.62 47	1.22 48	1.331 2	0.49 82	1.404 9	1.098 3	1.58 68
	100	BR	0.88 28	0.26 47	1.194 7	0.76 3	1.36 44	1.057	0.45 06	1.351 7	1.046 8	1.53 81
		GMBR	0.751 2	0.06 27	0.99 37	0.56 23	1.162 1	1.044 6	0.25 02	1.150 9	0.84 94	1.337 1
	250	BR	0.873 6	0.22 46	1.155 7	0.725 8	1.325 4	0.99 3	0.297 4	1.201	0.901	1.38 82
		GMBR	0.70 25	0.02 48	0.95 51	0.525 1	1.125 1	0.87 4	0.09 96	1.00 01	0.70 01	1.188 1
20%	30	BR	0.96 58	0.477 6	1.353 6	0.94 58	1.53 84	1.406 5	1.102 3	1.912 6	1.650 3	2.12 33
		GMBR	0.93 69	0.214 5	1.196 9	0.70 9	1.30 71	1.698 6	0.84 35	1.728 8	1.433 3	1.917 4
	50	BR	0.925 4	0.43 62	1.324 3	0.74 31	1.44 82	1.204 5	0.69 87	1.614 8	1.312 5	1.81
		GMBR	0.83 22	0.124 2	1.155 1	0.69 54	1.225 8	1.327 4	0.49 81	1.399 9	1.100 4	1.58 93
	100	BR	0.89 58	0.276 3	1.205 2	0.778 7	1.375 7	1.065 8	0.46 26	1.360 6	1.067 8	1.549 2
		GMBR	0.752 6	0.06 24	0.99 24	0.56 28	1.162 6	1.046 9	0.24 97	1.149 7	0.85 05	1.34 03
	250	BR	0.88 65	0.237 9	1.169 1	0.737 9	1.33 8	1.00 45	0.312 5	1.214 9	0.912 5	1.40 07
		GMBR	0.70 26	0.02 49	0.95 52	0.525 2	1.125 1	0.87 41	0.09 98	1.00 04	0.69 99	1.189 1

Table (1) show The GMBR method yields better results than BR because it achieves lower MSE and MAE values in most cases. Increasing the sample size reduces errors, which improves the accuracy of the model. Increasing the level of contamination increases the error values, but GMBR remains less affected than BR.

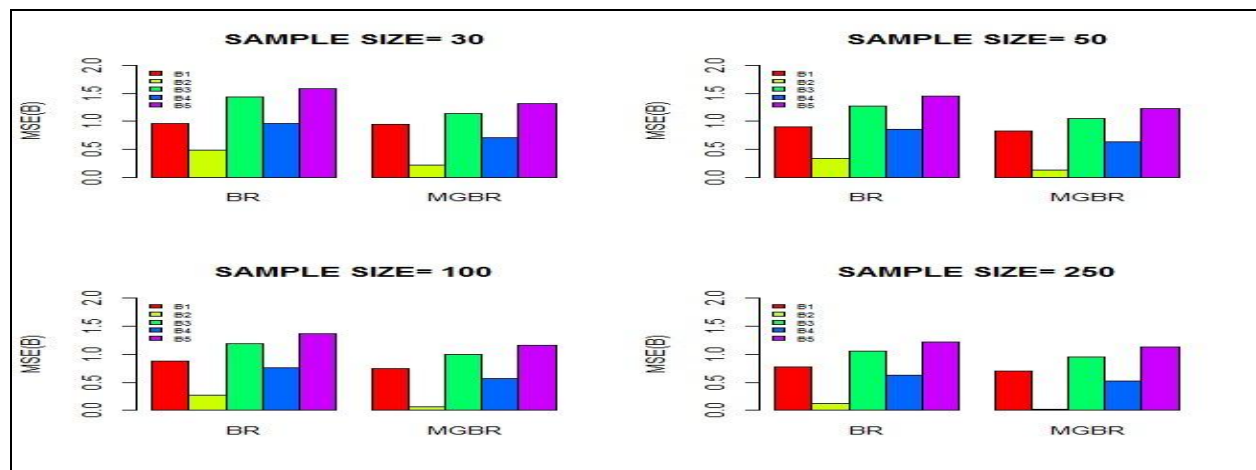


Figure (1): MSE ($\hat{\beta}$) for Different Sample Sizes at 5% Contamination in the First Scenario.

figure (1) show The GMBR method provides more accurate estimates and is less affected by contamination than the BR method. Increasing the sample size helps improve the accuracy of the estimate and reduce the MSE value for both methods, but GMBR remains the superior performer. Even at low contamination levels (5%), GMBR appears to be a more efficient method, making it a recommended choice when dealing with data that may contain anomalies or minor contamination.

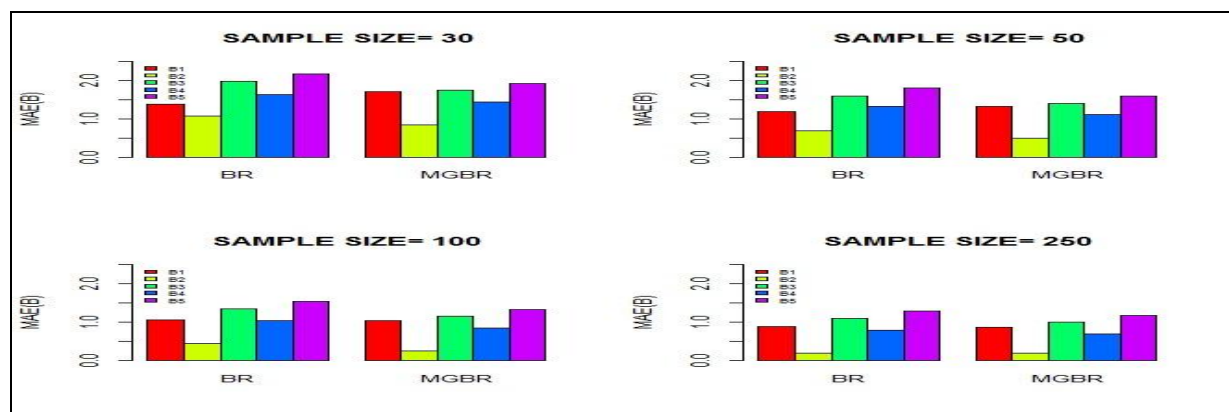


Figure (2): MAE ($\hat{\beta}$) for Different Sample Sizes at 5% Contamination in the First Scenario.

figure (2) show The GMBR method provides more accurate estimates and is less affected by contamination than the BR method. Increasing the sample size helps improve the accuracy of the estimate and reduces the MAE value for both methods, but GMBR remains the superior performer.

Even at low contamination levels (5%), GMBR appears to be a more efficient method, making it a recommended choice when dealing with data containing contamination or anomalies.

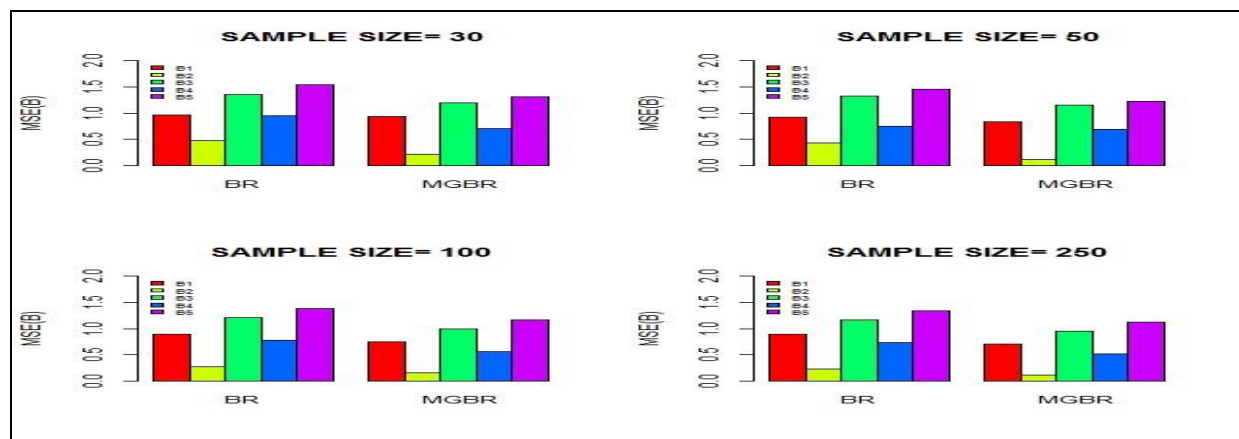


Figure (3): MSE ($\hat{\beta}$) for Different Sample Sizes at 20% Contamination in the First Scenario.

figure (3) show Increasing the contamination percentage (20%) significantly affects model performance, with higher MSE values for both methods, but the impact is greater for BR than for GMBR. GMBR provides more accurate estimates and is less susceptible to contamination, making it the preferred choice when dealing with contaminated data. As the sample size increases, MSE values decrease and the estimation accuracy improves, but GMBR still maintains an advantage over BR.

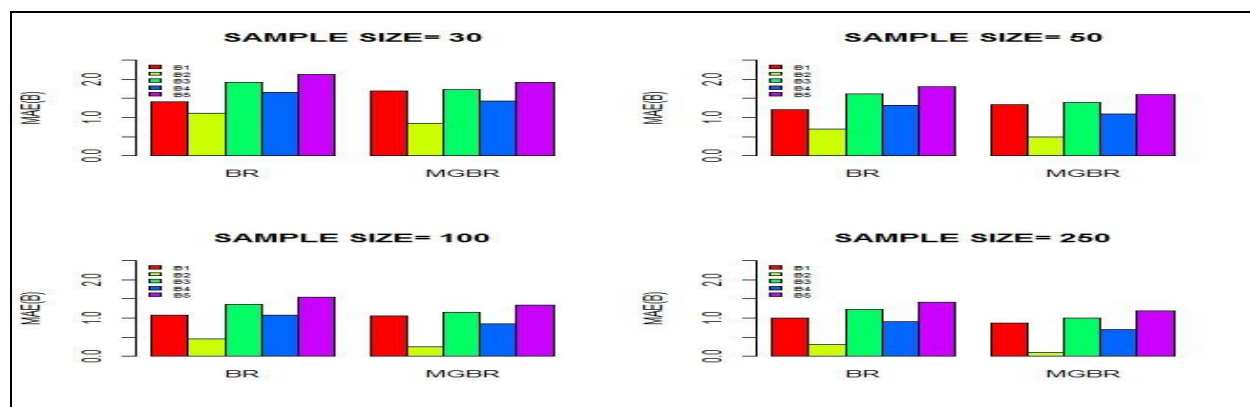


Figure (4): MAE ($\hat{\beta}$) for Different Sample Sizes at 20 % Contamination in the First Scenario.

figure (4) show Increasing the contamination percentage (20%) leads to higher errors in parameter estimation, making the model less accurate, but GMBR remains superior to BR. GMBR provides more stable estimates and is less affected by contamination, making it the preferred choice when dealing with contaminated data. As the sample size increases, MAE values decrease and estimation accuracy improves, but GMBR still maintains an advantage over BR, indicating that it is more robust against contaminated data.

Table (2): the MSE ($\hat{\beta}$) and MAE ($\hat{\beta}$) at Three Different Contamination Levels and different Sample Sizes in the First Scenario When the Precision Parameter is $\lambda=4$.

Con t.	Samp le	metho ds	MSE ($\hat{\beta}$)					MAE ($\hat{\beta}$)				
			$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$	$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$
5%	30	BR	4.84 51	2.418 7	2.247 9	1.368 6	1.845 4	5.941 3	3.597 1	3.572 1	2.542 7	3.03 09

		GMB R	1.273 2	1.505 3	1.298 3	0.42 51	0.94 5	2.90 6	2.758 8	2.66 07	1.672 5	2.210 2
	50	BR	3.26 37	2.20 34	2.00 89	1.098 8	1.670 9	3.939 5	2.935 4	2.86 24	1.798	2.43 43
		GMB R	0.761 7	1.338 4	1.125 1	0.24 47	0.78 52	1.743 7	2.09 83	1.987 6	0.98 64	1.562 4
	100	BR	2.06 52	2.06 34	1.842 1	0.98 04	1.509 9	2.40 31	2.44 58	2.319	1.353 5	1.90 86
		GMB R	0.37 5	1.212 9	0.99 45	0.125 8	0.65 58	0.86 54	1.599 4	1.478 5	0.501 1	1.056 1
	250	BR	1.321 3	1.984 7	1.767 3	0.89 8	1.427 8	1.462 2	2.145 3	2.02 64	1.047 9	1.602 4
		GMB R	0.14 69	1.137 9	0.92	0.05 01	0.58	0.34 27	1.299 7	1.179 9	0.20 01	0.755
10 %	30	BR	5.29 73	3.00 54	2.794 5	1.953 8	2.42 25	6.437 9	4.22 03	4.103 9	3.185	3.621 2
		GMB R	1.239 8	1.508 1	1.289 5	0.42 67	0.941 7	2.867 2	2.76 84	2.64 61	1.680 7	2.20 57
	50	BR	3.90 56	2.779 8	2.58 49	1.735 1	2.24 92	4.558 4	3.50 39	3.431 7	2.48 84	3.011 7
		GMB R	0.75 3	1.333 5	1.120 4	0.255 7	0.78 09	1.732 5	2.08 95	1.979 3	1.00 97	1.555 3
	100	BR	2.67 81	2.66 31	2.445 6	1.581 3	2.109 1	3.012 2	3.04 41	2.92 6	1.956 9	2.50 65
		GMB R	0.36 76	1.212 6	0.99 46	0.125 9	0.65 54	0.85 69	1.598 8	1.478 9	0.501 5	1.055 6
	250	BR	1.926 9	2.58 83	2.370 7	1.500 1	2.02 99	2.06 67	2.75	2.63 05	1.649 4	2.20 39
		GMB R	0.14 35	1.137 9	0.92 01	0.05	0.57 99	0.33 87	1.299 8	1.180 1	0.199 9	0.75 48
20 %	30	BR	5.74 66	3.255 6	3.00 73	2.131 4	2.67 03	6.815 6	4.50 74	4.30 87	3.32 31	3.901 2
		GMB R	1.268 1	1.514 4	1.287 1	0.414 4	0.95 02	2.89 98	2.781 8	2.64 21	1.656 9	2.22 36
	50	BR	4.126 2	3.03 95	2.80 9	1.927	2.46 44	4.781 1	3.80 76	3.66 35	2.65 97	3.22 86
		GMB R	0.72 93	1.341 2	1.121 6	0.247 8	0.779 1	1.705	2.105 6	1.982 5	0.99 49	1.552 5
	100	BR	2.89 35	2.89 78	2.677	1.803 9	2.33 65	3.231 7	3.28 65	3.161 1	2.174 7	2.735 3
		GMB	0.35	1.213	0.99	0.124	0.65	0.84	1.601	1.48	0.49	1.054

		R	73	7	53	6	5	48	3	04	91	9
	250	BR	2.161 1	2.818 5	2.60 08	1.730 3	2.25 93	2.29 95	2.98 06	2.861 1	1.88 02	2.43 32
		GMB R	0.141 7	1.138 1	0.92 01	0.05	0.57 98	0.33 66	1.30 02	1.180 2	0.2	0.75 47

Table (2) show When using a higher precision factor ($\lambda=4$), the MSE and MAE values are larger than in the first table ($\lambda=2$), but the GMBR method still provides better estimates compared to BR. Increasing the sample size reduces the error, which means the model's accuracy in estimating the parameters improves. Increasing the level of contamination increases the MSE and MAE values, but GMBR remains less affected by contamination than BR. Compared to the first table, we see that the higher precision factor ($\lambda=4$) slightly increases the error, but does not change the overall trend of the results.

Table(3): the MSE ($\hat{\beta}$) and MAE ($\hat{\beta}$) at Three Different Contamination Levels and different Sample Sizes in the Second Scenario When the Precision Parameter is $\lambda=2$.

Co nt.	Sa mp le	met hod s	MSE ($\hat{\beta}$)							MAE ($\hat{\beta}$)						
			$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$	$\hat{\beta}_6$	$\hat{\beta}_7$	$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$	$\hat{\beta}_6$	$\hat{\beta}_7$
5 %	30	BR	9.0 501	2.1 77	1.5 70 6	2.7 231	1.4 94 8	2.4 74 9	1.5 98	5.9 20 5	2.8 718	2.3 11	3.4 587	2.2 52 3	3.3 38	2.3 38 6
		GM BR	6.5 251	1.3 913	0.8 321	1.9 46	0.8 20 3	1.7 119	0.8 38 2	4.6 617	2.1 937	1.6 60 6	2.7 627	1.6 49 9	2.6 50 8	1.6 667
	50	BR	10. 881 6	2.2 28 6	1.7 06 3	2.2 45 6	1.71 75	2.6 174	1.7 04	7.0 42 2	3.1 541	2.6 47	3.7 86	2.6 67 8	3.6 74	2.6 617
		GM BR	7.8 243	1.5 45 8	0.9 99 8	2.1 114	1.0 00 4	1.8 779	0.9 98 8	5.5 93 2	2.5 182	1.9 96 6	3.0 96 6	1.9 977	2.9 84 9	1.9 96 5
	100	BR	5.6 941	1.7 63 6	1.1 87	2.2 58	1.1 83 6	2.0 37 3	1.17 98	2.6 34 6	2.2 279	1.6 771	2.7 36	1.6 731	2.6 36	1.6 68 4
		GM BR	3.8 863	1.0 621	0.5 014	1.5 951	0.5 013	1.3 67 9	0.5 00 2	2.7 87 6	1.5 36 6	1.0 00 9	2.0 82 6	1.7 00 7	1.9 773	0.9 99 6
	25 0	BR	2.6 999	1.4 26 8	0.8 691	1.9 72 6	0.8 70 5	1.7 42 3	0.8 725	1.9 437	1.6 00 4	1.0 677	2.1 59 3	1.0 69 3	2.0 50 8	1.0 70 9
		GM BR	1.55 53	0.7 553	0.1 99 4	1.3 00 3	0.1 99 9	1.0 70 2	0.2 00 5	1.11 54	0.9 30 2	0.3 99 3	1.4 88 2	0.3 99 8	1.3 80 1	0.4 00 4
10	30	BR	17.3	3.4	2.8	3.9	2.8	3.7	2.9	11.	4.8	4.3	5.4	4.3	5.3	4.4

%			64	02 2	05 3	48 4	48 2	28 7	24 5	391 2	82 5	26 6	28	63 8	74 2	106	
		GM BR	12. 944	2.2 277	1.6 67	2.7 67 4	1.6 746	2.5 56 8	1.6 88 6	9.2 85 3	3.8 60 9	3.3 25 2	4.4 13	3.3 33 4	4.3 26 5	3.3 45 9	
	50	BR	11.0 846	2.6 53 6	2.0 556	3.1 86	2.0 587	2.8 89 5	2.0 554	7.3 40 3	3.5 88 2	3.0 18 8	4.1 29 5	3.0 17	3.9 591	3.0 166	
		GM BR	7.7 851	1.5 65	1.0 012	2.1 06 2	1.0 03 3	1.8 63 4	1.0 00 5	5.5 78 8	2.5 38	1.9 99 2	3.0 917	2.0 011	2.9 712	1.9 98 5	
	100	BR	6.1 086	2.0 59 8	1.5 23	2.6 318	1.5 34 9	2.3 96 9	1.5 47	4.2 04 4	2.5 261	2.0 156	3.1 116	2.0 27 6	2.9 98 3	2.0 39 4	
		GM BR	3.8 926	1.0 5	0.4 991	1.6 01 3	0.5 02 2	1.3 69 5	0.5 05 6	2.7 89 8	1.5 24 6	0.9 98 7	2.0 88 9	1.0 01 8	1.9 79	1.0 05 2	
	25 0	BR	3.0 652	1.7 76	1.2 215	2.3 194	1.2 213	2.0 92 3	1.2 201	2.2 98 4	1.9 5	1.4 20 4	2.5 06 3	1.4 20 3	2.4 012	1.4 191	
		GM BR	1.55 66	0.2	0.2 00 2	1.2 99 8	0.2	1.0 70 4	0.1 99 8	1.11 59	0.9 29 9	0.4 00 1	1.4 877	0.3 99 9	1.3 80 4	0.3 997	
	20 %	30	BR	17.4 993	3.0 613	3.1 614	5.7 413	3.9 69 4	3.0 351	4.5 36 8	11. 615 7	6.0 06 8	5.0 281	4.6 05 8	5.5 30 5	7.3 05 8	4.6 86 6
			GM BR	12. 89	1.6 64 3	1.6 96 6	4.3 67 9	2.5 70 4	1.6 66 2	3.0 80 3	9.2 655	4.7 247	3.6 57 8	3.3 29 3	4.2 318	6.0 27 9	3.3 40 6
		50	BR	11.3 147	2.3 516	2.3 371	5.0 07	3.1 967	2.3 87 4	3.6 92 9	7.6 02 4	4.7 251	3.6 091	3.2 774	4.1 70 6	6.0 579	3.2 58 8
			GM BR	7.77 34	1.0 061	1.0 051	3.6 95 4	1.8 94 9	1.0 145	2.3 92 3	5.5 745	3.4 04 7	2.3 03 5	1.9 93 8	2.8 93 5	4.7 13	1.9 90 5
100		BR	6.3 397	1.7 99 9	1.8 106	4.5 167	2.7 05 2	1.8 02 5	3.2 018	4.4 69 3	3.6 94 9	2.6 06 2	2.3 09 7	3.1 99 8	4.9 96 3	2.2 96 3	
		GM BR	3.8 755	0.4 99	0.5 013	3.2 02 4	1.4 01	0.4 99	1.8 99 6	2.7 83 7	2.3 98 7	1.3 01	1.0 021	1.9 00 7	3.6 98 7	0.9 99 3	
25 0		BR	3.3 305	1.5 03 3	1.5 018	4.1 98 7	2.4 02 4	1.5 00 1	2.8 98 7	2.5 735	3.1 02 5	2.0 01	1.6 979	2.6 016	4.3 99 3	1.6 97 8	

		GM BR	1.5 472	0.2 00 7	0.2 00 2	2.8 99 6	1.1 00 3	0.2	1.5 99 4	1.11 25	1.8 00 7	0.7 00 2	0.3 99 5	1.3 00 3	3.1	0.3 99 4
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Table (3) show In the second scenario (when $\lambda=2$), the GMBR method still outperforms BR, achieving lower MSE and MAE values in most cases. Increasing the sample size significantly reduces the error, which means the model's accuracy in estimating parameters improves. Increasing the pollution level significantly increases the MSE and MAE values, but GMBR remains less affected by pollution than BR. Compared to the previous tables, the second scenario appears to contain more variables ($\hat{\beta}_1$ to $\hat{\beta}_7$), making estimation more complex, but GMBR still performs better than BR.

Table(4): the MSE ($\hat{\beta}$) and MAE ($\hat{\beta}$) at Three Different Contamination Levels and different Sample Sizes in the Second Scenario When the Precision Parameter is $\lambda=4$.

Co nt.	Sa mp le	met hod s	MSE ($\hat{\beta}$)							MAE ($\hat{\beta}$)						
			$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$	$\hat{\beta}_6$	$\hat{\beta}_7$	$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$	$\hat{\beta}_6$	$\hat{\beta}_7$
5 %	30	BR	6.1 25 3	3.2 93	3.1 901	3.6 73 8	2.6 69 9	3.5 169	4.6 59 6	7.1 031	4.9 06 9	4.9 88 3	5.2 091	4.1 88 2	5.0 55 4	5.9 07 9
		GM BR	4.4 68 7	2.3 49 9	2.2 728	2.7 731	1.6 912	2.5 78 2	3.8 45 8	5.4 752	4.0 98 3	4.2 016	4.4 20 9	3.3 495	4.2 65 6	5.2 05 6
	50	BR	5.1 78 8	2.5 105	2.4 45	2.9 48 5	1.8 197	2.6 66 5	4.0 28	5.0 27 6	3.5 59 6	3.6 76 6	3.8 99 3	2.7 84 4	3.6 69 6	4.6 90 4
		GM BR	3.9 27	1.6 753	1.6 06 8	2.1 07 7	1.0 00 1	1.8 91	3.2 012	3.3 00 5	2.7 63	2.8 747	3.0 93 6	1.9 981	2.9 195	3.8 99
	100	BR	5.3 02 8	1.9 551	1.9 00 2	2.3 87 4	1.3 09 8	2.1 94 4	3.4 93 6	4.3 72	2.5 379	2.6 62 5	2.8 68 2	1.8 013	2.7 165	3.6 85 8
		GM BR	3.9 341	1.1 68 4	1.1 014	1.5 98 2	0.5 03 8	1.3 99 8	2.6 99 9	3.1 412	1.7 58	1.8 70 9	2.0 85 8	1.0 03 3	1.9 29 3	2.8 99 4
	250	BR	6.0 317	1.7 09 5	1.6 357	2.1 39 5	1.0 39 6	1.9 39 5	3.2 421	3.0 776	2.0 48 3	2.1 54 2	2.3 761	1.2 88 5	2.2 183	3.1 90 7
		GM BR	2.4 66 4	0.9 20 3	0.8 49 4	1.3 501	0.2 501	1.1 50 2	2.4 50 5	1.5 70 4	1.2 60 2	1.3 69 3	1.5 88	0.5	1.4 301	2.4 00 4
10 %	30	BR	5.5 661	1.2 24 4	1.1 631	1.6 89 9	0.5 757	1.4 58 2	2.7 56 3	2.5 90 5	1.5 49 4	1.6 657	1.9 112	0.8 07 8	1.7 193	2.6 9
		GM	2.4	0.9	0.8	1.3	0.2	1.1	2.4	1.5	1.2	1.3	1.5	0.5	1.4	2.3

		BR	63 9	179	5	56 6	53	48 5	48 7	69 3	57	69	93 6	019	27 4	977
	50	BR	6.3 68 2	2.0 251	1.9 63 4	2.4 56 3	1.3 617	2.2 48 7	3.5 475	3.3 92	2.3 58 2	2.4 77	2.6 881	1.6 03 6	2.5 216	3.4 901
		GM BR	2.4 66 9	0.9 20 3	0.8 522	1.3 50 5	0.2 517	1.1 48 7	2.4 481	1.5 70 4	1.2 6	1.3 718	1.5 881	0.5 013	1.4 28 3	2.3 977
	100	BR	6.3 65 6	2.0 163	1.9 49 3	2.4 59	1.3 472	2.2 50 3	3.5 579	3.3 92 8	2.3 531	2.4 66 4	2.6 941	1.5 94 4	2.5 277	3.5 051
		GM BR	2.4 68 2	0.9 188	0.8 495	1.3 519	0.2 48 6	1.1 5	2.4 518	1.5 71	1.2 58 6	1.3 69 3	1.5 89 8	0.4 98 5	1.4 29 8	2.4 016
	250	BR	6.3 50 3	2.0 23	1.9 48 3	2.4 49 2	1.3 534	2.2 48 7	3.5 46 3	3.3 90 2	2.3 619	2.4 67 4	2.6 86 3	1.6 02 3	2.5 277	3.4 952
		GM BR	2.4 63 6	0.9 20 8	0.8 497	1.3 49 8	0.2 50 9	1.1 49 7	2.4 49 2	1.5 69 5	1.2 60 7	1.3 69 6	1.5 87 8	0.5 00 8	1.4 29 7	2.3 991
20 %	30	BR	5.5 92 3	1.2 535	1.1 83 5	1.6 79	0.5 90 8	1.4 735	2.7 753	2.6 10 8	1.5 83 8	1.6 94 7	1.9 04 9	0.8 28 4	1.7 411	2.7 144
		GM BR	2.4 63 5	0.9 213	0.8 518	1.3 49 7	0.2 527	1.1 48 6	2.4 49 5	1.5 691	1.2 60 7	1.3 713	1.5 87	0.5 02	1.4 27 9	2.3 98 8
	50	BR	5.7 01 8	1.3 515	1.2 82 2	1.7 88 8	0.6 857	1.5 88 6	2.8 791	2.7 23	1.6 87 2	1.7 97 8	2.0 227	0.9 30 5	1.8 641	2.8 24
		GM BR	2.4 66 2	0.9 195	0.8 49 6	1.3 512	0.2 50 3	1.1 514	2.4 49	1.5 70 2	1.2 59 3	1.3 69 3	1.5 89	0.5	1.4 311	2.3 98 7
	100	BR	5.6 78 3	1.3 53	1.2 80 5	1.7 88 8	0.6 78 2	1.5 816	2.8 81	2.7 195	1.6 90 8	1.7 98 6	2.0 24 6	0.9 26 2	1.8 597	2.8 29
		GM BR	2.4 62 8	0.9 20 4	0.8 49 8	1.3 517	0.2 491	1.1 49 9	2.4 5	1.5 69 2	1.2 60 2	1.3 69 7	1.5 89 6	0.4 99	1.4 29 8	2.3 99 8
	250	BR	6.1 778	1.8 64 6	1.7 96 2	2.2 94 2	1.1 979	2.0 93 3	3.3 92 3	3.2 312	2.2 03 8	2.3 155	2.5 315	1.4 471	2.3 72 6	3.3 415
		GM BR	2.4 56 9	0.9 20 2	0.8 50 4	1.3 501	0.2 507	1.1 49 5	2.4 49 3	1.5 67 4	1.2 60 2	1.3 70 3	1.5 881	0.5 00 7	1.4 29 5	2.3 99 2

Table (4) show When using a higher precision factor ($\lambda=4$), the MSE and MAE values become larger than in Table 3 ($\lambda=2$), but GMBR still provides better estimates than BR. Increasing the sample size significantly reduces the error, which means the model's accuracy in estimating parameters improves. Increasing the level of contamination increases the MSE and MAE values, but GMBR remains less affected by contamination than BR. Compared to the previous tables, we see that a higher precision factor ($\lambda=4$) increases the error, but it does not change the overall trend of the results, as GMBR remains the most accurate and stable method.

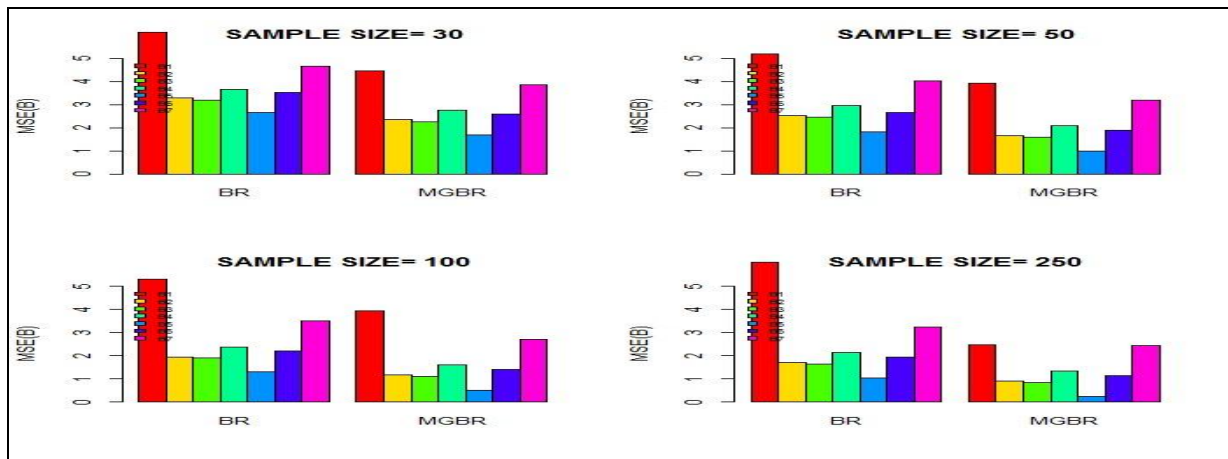


Figure (5): MSE ($\hat{\beta}$) for Different Sample Sizes at 5% Contamination in the Second Scenario.

figure (5) show The GMBr method provides more accurate estimates and is less susceptible to contamination than the BR method, even at low contamination levels (5%). Increasing the sample size significantly improves model accuracy and reduces the MSE value, but GMBr remains superior in all cases. The GMBr method is a better choice, especially when dealing with partially contaminated data or when using small samples.

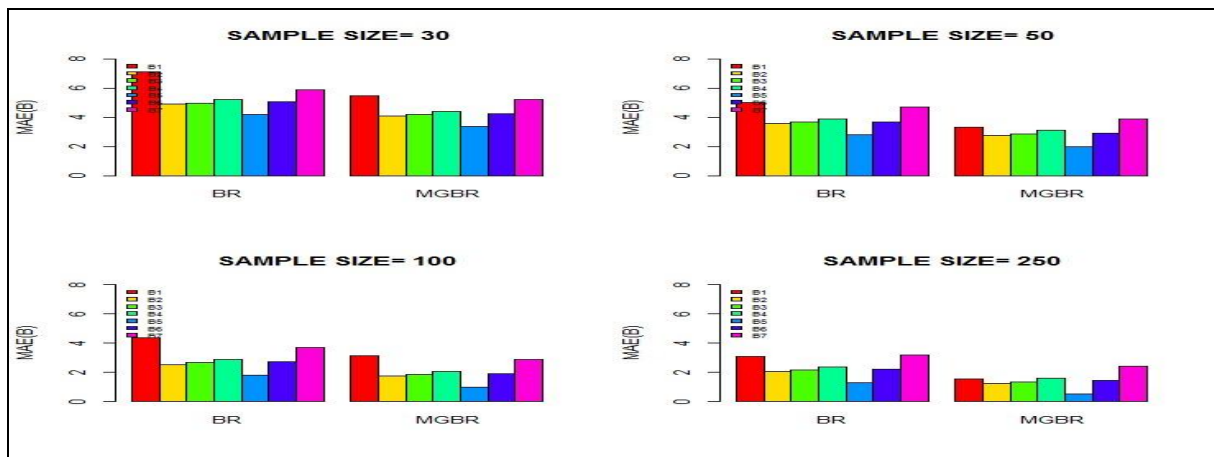


Figure (6): MAE ($\hat{\beta}$) for Different Sample Sizes at 5 % Contamination in the Second Scenario.

figure (6) show The GMBr method provides more accurate estimates and is less affected by contamination than the BR method, even at low contamination levels (5%). Increasing the sample size improves model accuracy and reduces the MAE value, but GMBr remains the best performer in all cases. The GMBr method is an ideal choice when stable and accurate estimates are needed, especially with partially contaminated data or small samples.

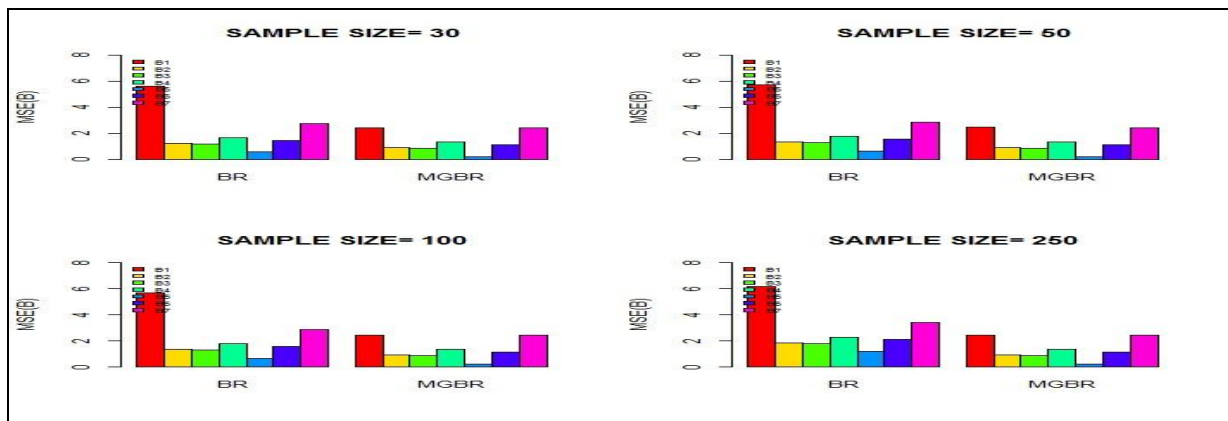


Figure (7): MSE ($\hat{\beta}$) for Different Sample Sizes at 20% Contamination in the Second Scenario.

figure (7) show Increasing the contamination percentage (20%) significantly affects model performance, with higher MSE values for both methods, but the effect is greater for BR than for GMBR. GMBR provides more accurate estimates and is less susceptible to contamination, making it the preferred choice when dealing with contaminated data. As the sample size increases, MSE values decrease and the estimation accuracy improves, but GMBR still maintains an advantage over BR, indicating that it is more robust against contaminated data.

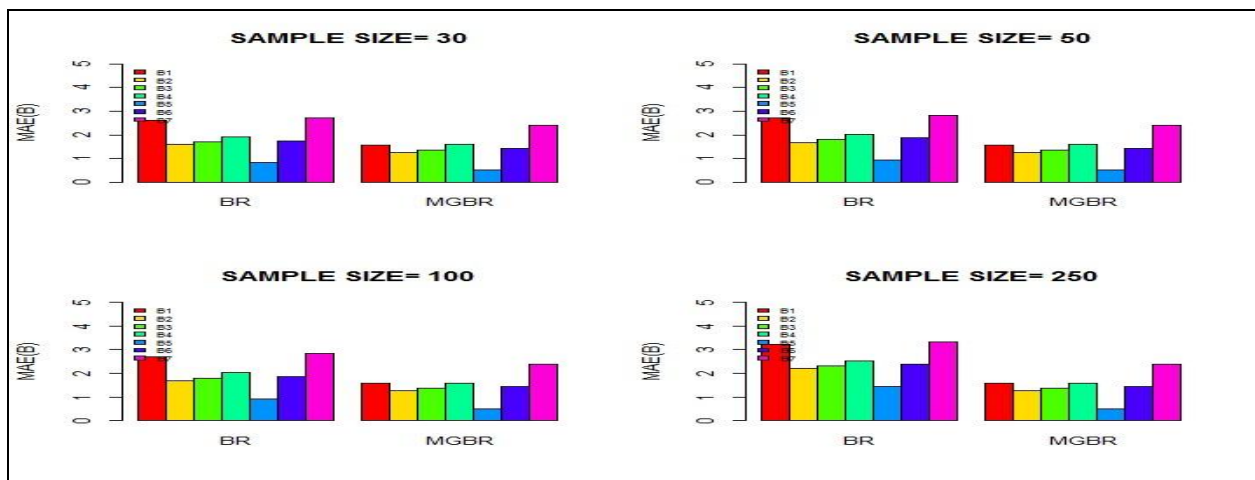


Figure (8): MAE ($\hat{\beta}$) for Different Sample Sizes at 20 % Contamination in the Second Scenario.

figure (8) show Increasing the contamination percentage (20%) significantly affects the estimation accuracy, as MAE values increase for both methods, but the effect is greater for BR than for GMBR.

GMBR provides more accurate estimates and is less susceptible to contamination, making it the preferred choice when dealing with contaminated data. As the sample size increases, MAE values decrease and estimation accuracy improves, but GMBR still maintains an advantage over BR, indicating that it is more robust against contaminated data.

Real Data

The Gasoline Yield dataset, originally gathered by Prater in 1956, focuses on the proportion of crude oil remaining after distillation and fractionation. Atkinson (1985) later analyzed the data using a linear regression model and observed that the error distribution appeared slightly asymmetrical, resulting in unusually large and small residuals. For our study, we introduced contamination by altering the first observations of the explanatory variables by 10%.

Dataset Description

The Gasoline Yield dataset contains **32 observations** of **6 variables** related to gasoline reduction:

Variable Description

yield Proportion of crude oil converted into gasoline (response variable)

gravity API gravity of crude oil

pressure Vapor pressure of crude oil

temp10 Temperature at which 10% of crude oil has vaporized

temp Temperature at which 50% of crude oil has vaporized

temp90 Temperature at which 90% of crude oil has vaporized

Table(5): the coefficient estimated by BR and GMBR methods

methods	$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$	$\hat{\beta}_6$
BR	-1.50568	1.69601	20.35790	-1.64932	2.87122	-23.13934
GMBR	-0.19659	0.58397	1.67494	-0.36112	0.40733	-2.28240

Table (5) show The GMBR method provides more stable and less extreme estimates than BR, making it more accurate and reliable. The BR method may be affected by outliers or contamination, leading to unstable or highly extreme estimates. GMBR is preferred when dealing with data that may contain significant noise or variance, as it provides more conservative and stable estimates.

Table (6): the MSE and MAE to the model that estimated by BR and GMBR.

	MSE	MAE
BR	0.30175	0.54891
MGBR	0.14614	0.29730

Table (6) show The GMBR method achieves lower MSE and MAE values compared to the BR method. This suggests that GMBR is more accurate and stable in model estimation.

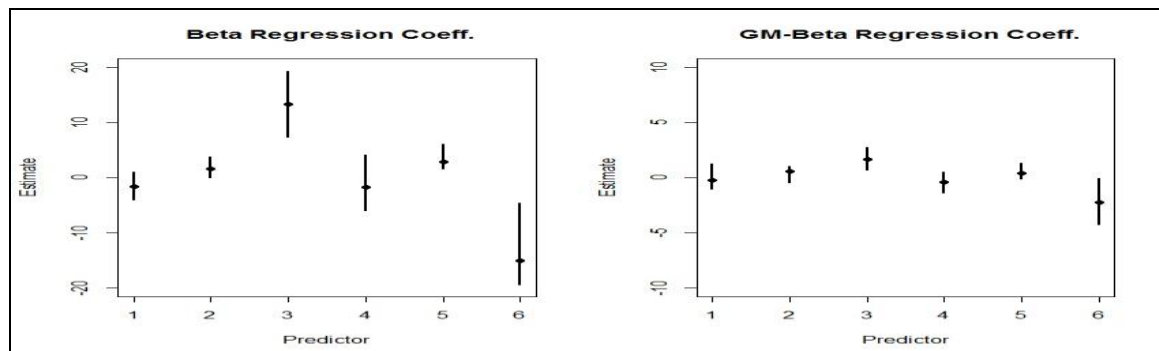


Figure (9): Confidence Intervals for the Estimated Coefficients Using BR and GMBR Methods.

figure (9) show The GMBR method achieves more stable estimates and is less susceptible to dispersion than the BR method, as demonstrated by the narrower confidence intervals. The BR method exhibits

greater variance in parameter estimates, suggesting it may be less accurate, especially when dealing with data containing anomalies or contamination. The GMBR method is preferred when the goal is to obtain more reliable estimates that are less sensitive to contaminated data.

Conclusions:

This study examines beta regression analysis and its use in statistical modeling when the response values are between 0 and 1, focusing on high leverage points of parameter estimates. The maximum weighted estimation (MLE) method was used to estimate model parameters, but this method is sensitive to data anomalies, which can lead to skewed conclusions. Therefore, the generalized M-estimator (GM-estimator) method was proposed to reduce the influence of high leverage points and enhance model accuracy. The performance of the GMBR method was compared with the traditional beta regression (MLBr) method through simulation experiments and real data using criteria such as the mean squared error (MSE) and mean absolute error (MAE). The Generalized M-estimator (GMBR) method demonstrated superior performance over the traditional beta regression (MLBr) method, providing more accurate and stable estimates by reducing the mean squared error (MSE) and mean absolute error (MAE) across most sample sizes and contamination levels.

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